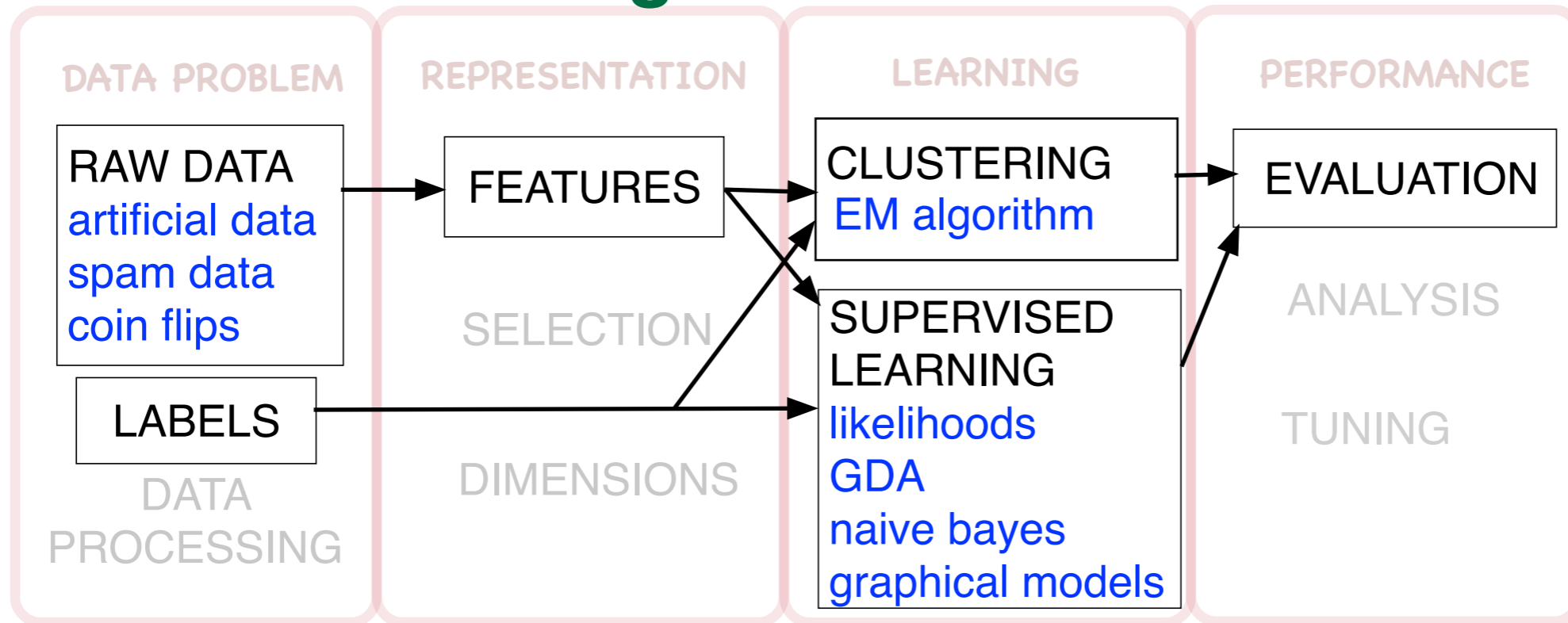


Gaussian Discriminant Analysis

material thanks to Andrew Ng @Stanford

Course Map / module3

module 3: generative methods



- Gaussian Discriminant Analysis

Density Estimation Problem

- $P(y|x) = P(y|x^1, x^2, \dots, x^d)$ joint $(d+1)$ -dim distribution
- ... actually we cannot estimate this joint
- if each feature has 10 buckets, and we have 100 features (very reasonable assumptions)
- then the joint distribution has 10^{100} cells - impossible

how to get around estimating the joint $P(x^1, x^2, \dots, x^d | y)$?

- SOLUTION: model/restrict the joint, instead of estimating any possible such joint distribution
 - fore example with a well known parametrized form
 - such as multi-dim gaussian distribution
 - estimate the parameters of the imposed model
- called **Gaussian Discriminant Analysis** (when the model imposed is gaussian)
 - easy to implement due to math tools facilitating gaussian parameters estimation (mean, covariance)
 - multidim implies “covariance” matrix instead of simple variance
 - doesnt fit data in many cases

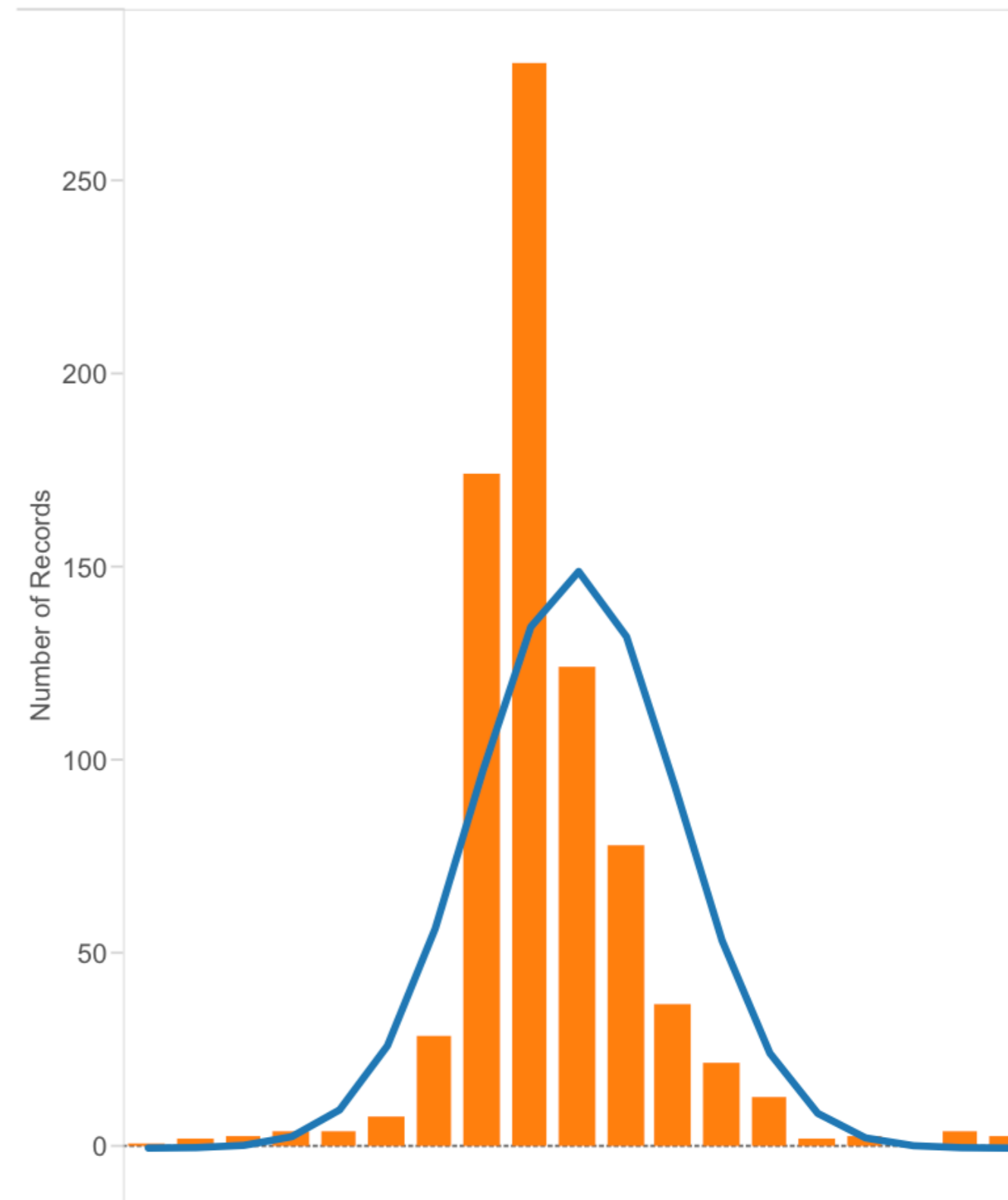
Gaussian Fit

- Idea: fit a parametrized distribution to histogram (density or counts)
- The gaussian (normal) density is controlled by mean and variance

$$P(x|\mu, \sigma^2) = \text{normal}(x, \mu, \sigma^2) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

- the best fit is the one that maximizes likelihood of the data

$$\log L = \log \prod_{i=1}^m P(x_i|\mu, \sigma^2) = \sum_{i=1}^m \log P(x_i|\mu, \sigma^2)$$



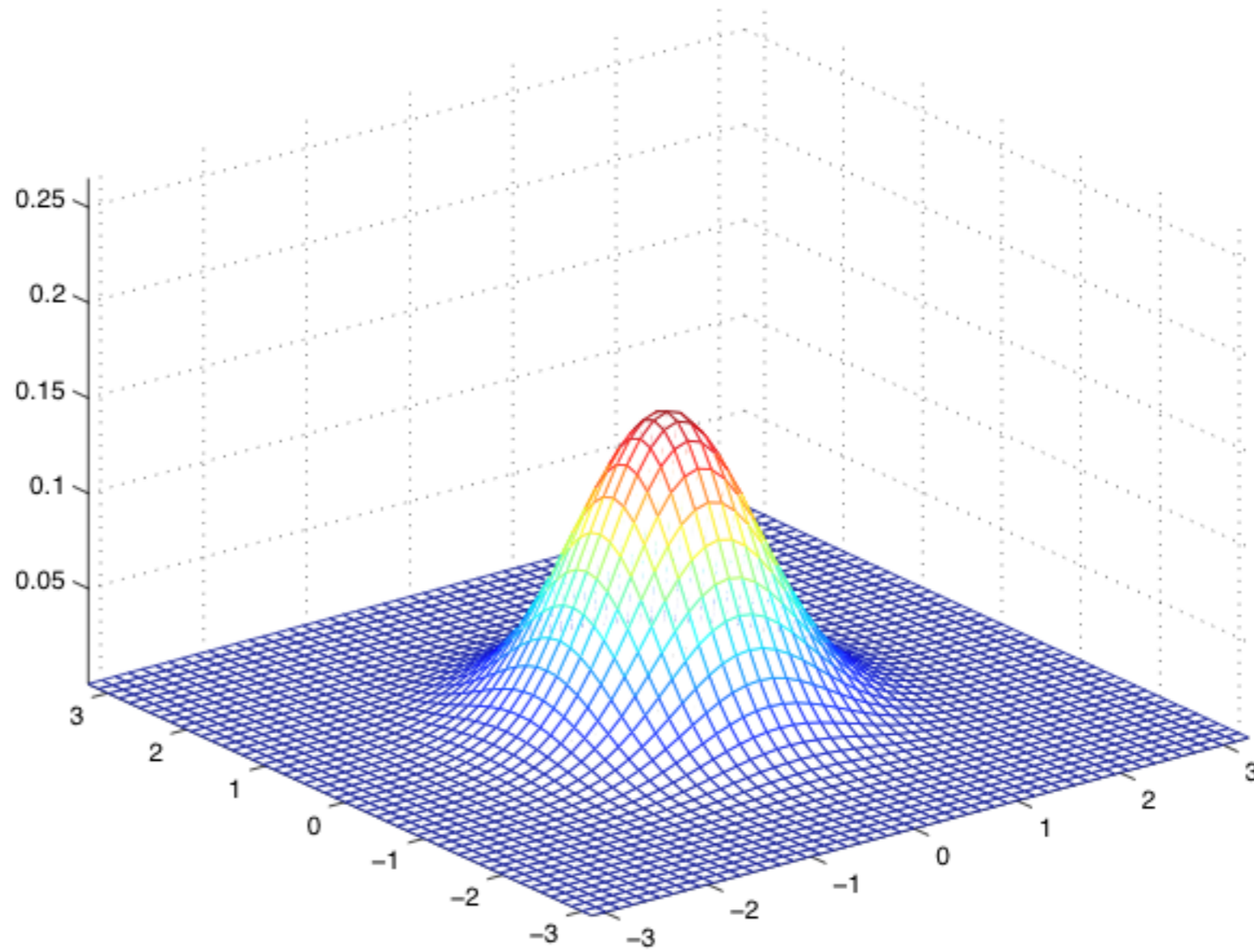
Lets impose a nice probabilistic model

- Multi-variate normal $\theta = (\mu, \Sigma)$
distribution

$$p(x; \mu, \Sigma) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left(-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu) \right)$$

- plotted Σ =identity (or independent variables)

$$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

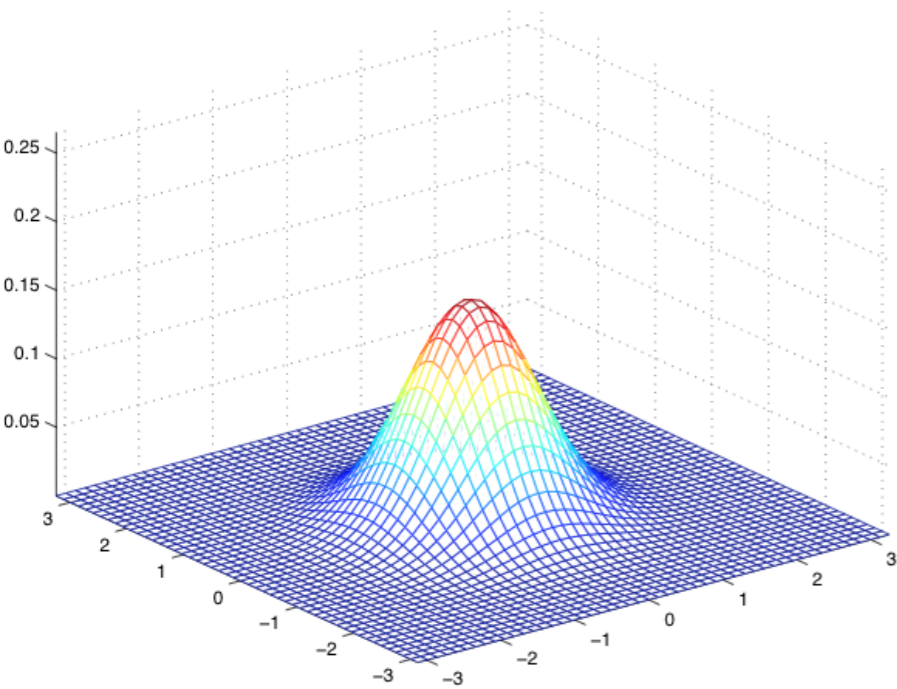


Lets impose a nice probabilistic model

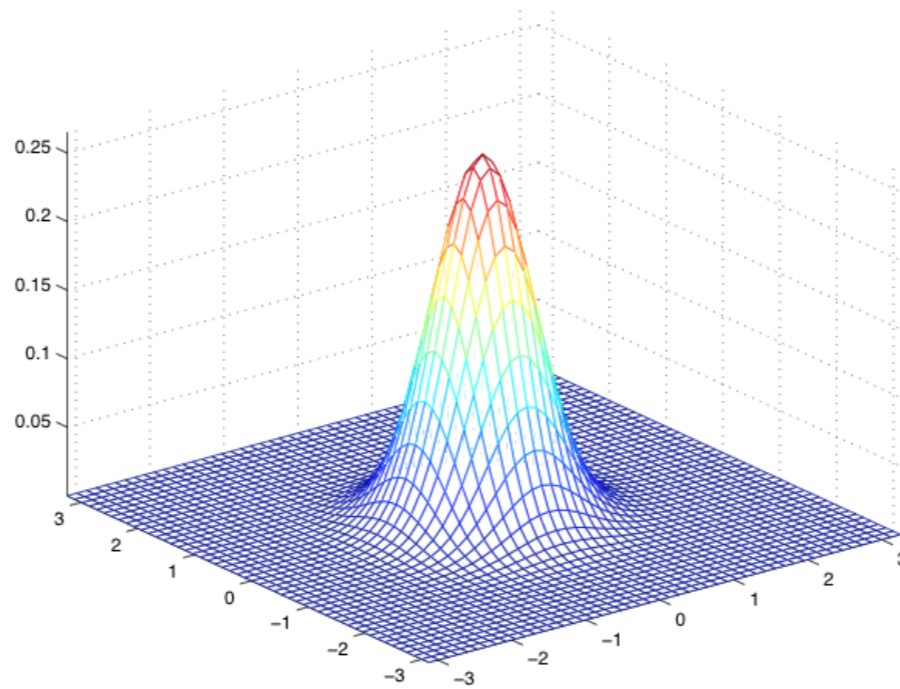
- Multi-variate normal distribution

$$p(x; \mu, \Sigma) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left(-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu) \right)$$

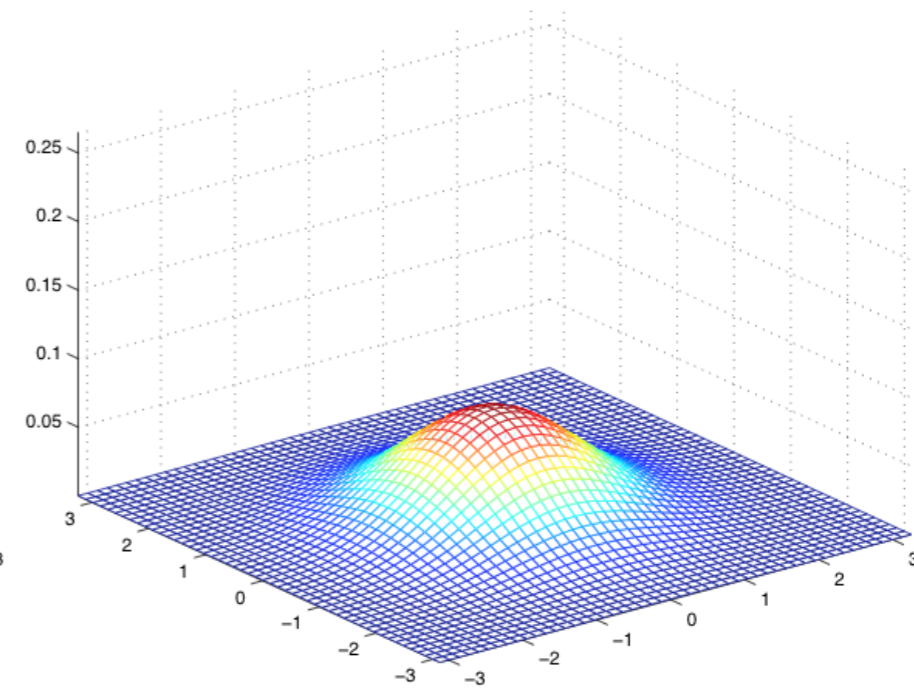
- plotted Σ =variance only or independent variables



$$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 0.6 & 0 \\ 0 & 0.6 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}$$

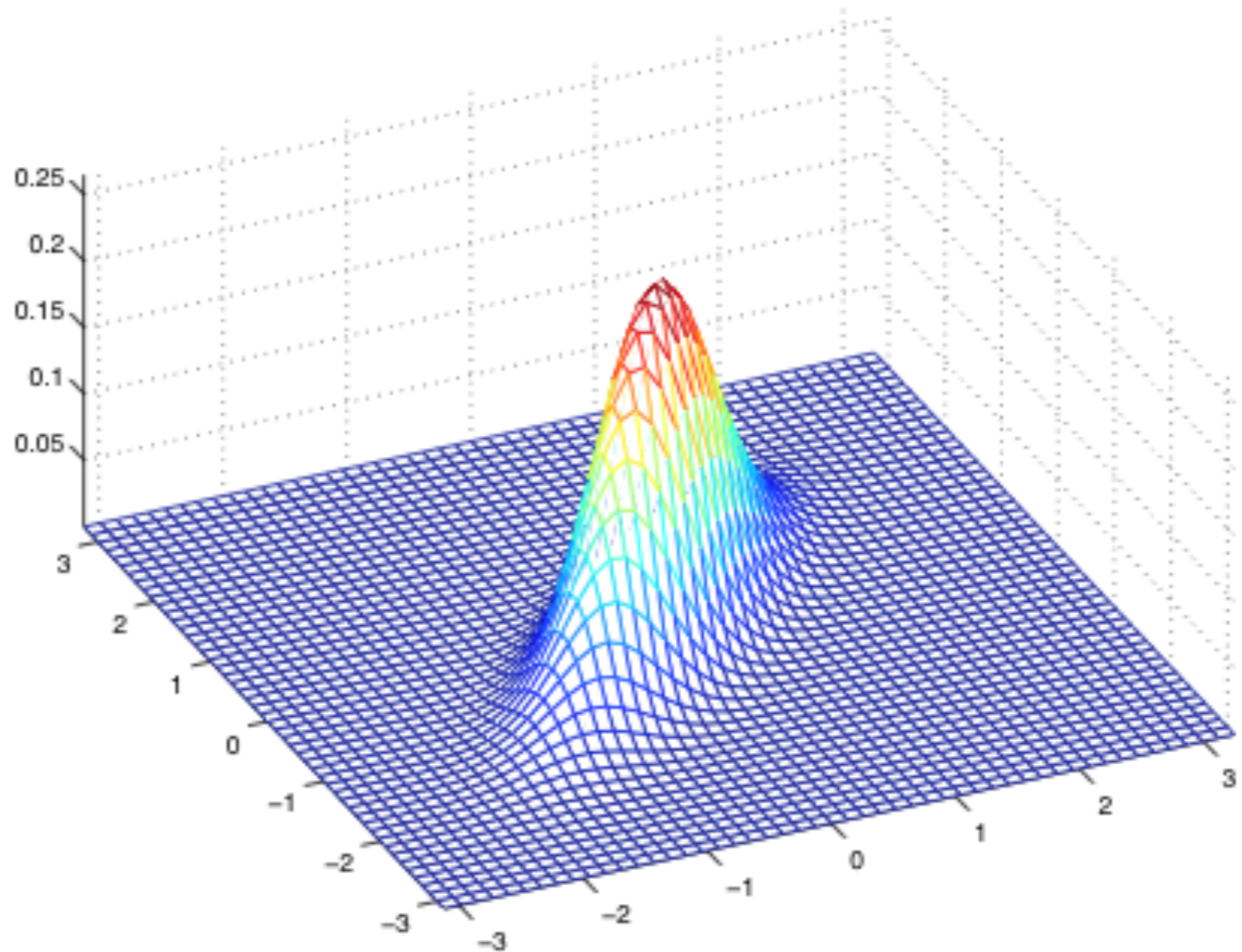
Lets impose a nice probabilistic model

- Multi-variate normal distribution

$$p(x; \mu, \Sigma) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left(-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu) \right)$$

- plotted $\Sigma \neq \text{identity}$
- dependent variables

$$\Sigma = \begin{bmatrix} 1 & 0.8 \\ 0.8 & 1 \end{bmatrix}$$

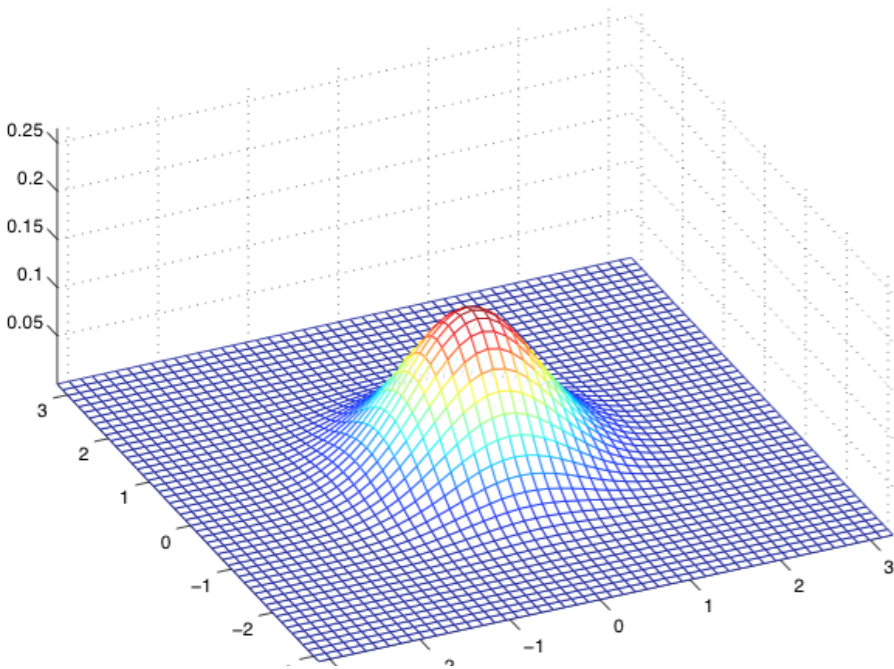


Lets impose a nice probabilistic model

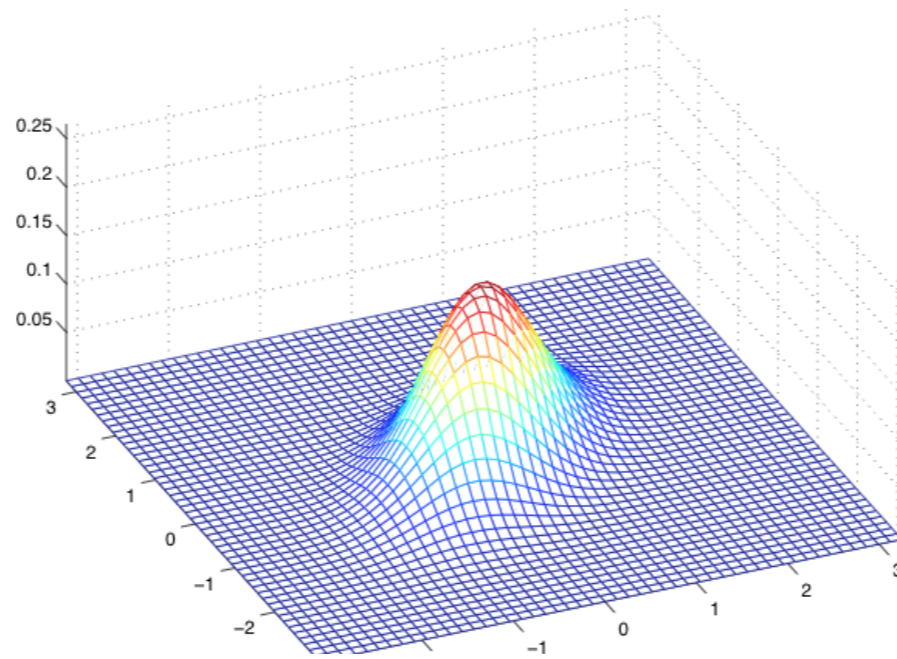
- Multi-variate normal distribution

$$p(x; \mu, \Sigma) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left(-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu) \right)$$

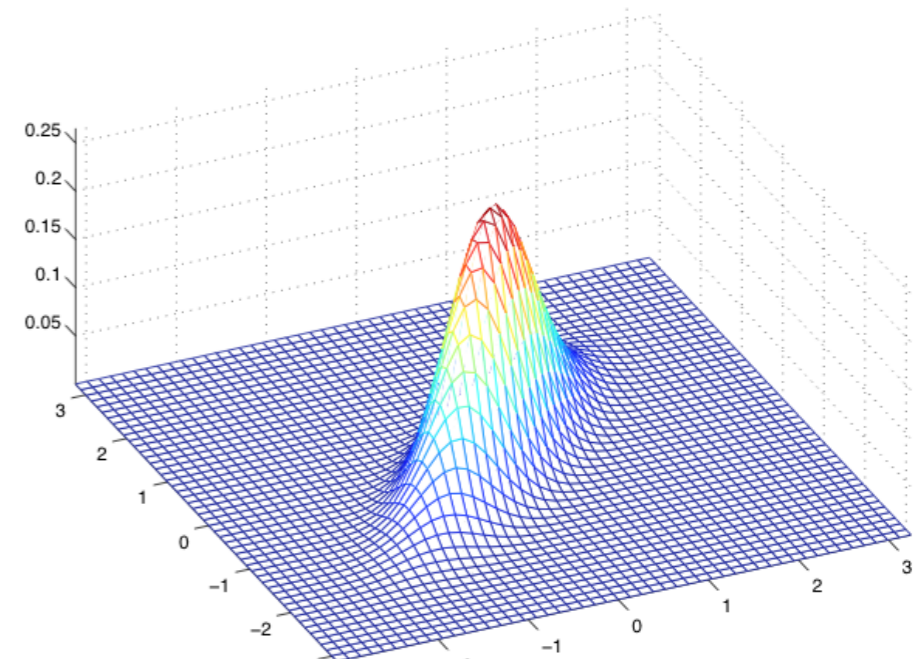
- $\Sigma \neq \text{identity} \Rightarrow$ dependent variables



$$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 1 & 0.8 \\ 0.8 & 1 \end{bmatrix}$$

GDA Setup

- multi normal density estimation for each y (common Σ)

$$p(x|y = 0) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left(-\frac{1}{2} (x - \mu_0)^T \Sigma^{-1} (x - \mu_0) \right)$$

$$p(x|y = 1) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left(-\frac{1}{2} (x - \mu_1)^T \Sigma^{-1} (x - \mu_1) \right)$$

- log likelihood

$$\begin{aligned} \ell(\phi, \mu_0, \mu_1, \Sigma) &= \log \prod_{i=1}^m p(x^{(i)}, y^{(i)}; \phi, \mu_0, \mu_1, \Sigma) \\ &= \log \prod_{i=1}^m p(x^{(i)} | y^{(i)}; \mu_0, \mu_1, \Sigma) p(y^{(i)}; \phi) \end{aligned}$$

GDA parameter solution

- max likelihood for GDA has close form solution!
- can be derived using differentials
 - estimate mean for each class
 - estimate covariance for entire training set
 - or separately for each class
 - no need for Gradient Descent or other optimizers

$$\phi = \frac{1}{m} \sum_{i=1}^m 1\{y^{(i)} = 1\}$$

$$\mu_0 = \frac{\sum_{i=1}^m 1\{y^{(i)} = 0\} x^{(i)}}{\sum_{i=1}^m 1\{y^{(i)} = 0\}}$$

$$\mu_1 = \frac{\sum_{i=1}^m 1\{y^{(i)} = 1\} x^{(i)}}{\sum_{i=1}^m 1\{y^{(i)} = 1\}}$$

$$\Sigma = \frac{1}{m} \sum_{i=1}^m (x^{(i)} - \mu_{y^{(i)}})(x^{(i)} - \mu_{y^{(i)}})^T$$

GDA visual classification

- if common Σ , the two gaussians are identical except for the mean
- the separation is a line of equidistant points to the two means

